

## Particle filtering for high-dimensional nonlinear systems: An introduction and new ideas

#### Jon Poterjoy

Department of Meteorology The Pennsylvania State University

Thursday 12th December, 2013



Our current state-of-the-art data assimilation systems (e.g., 4DVar, EnKF, and hybrid methods) operate under the assumptions of linear model dynamics and Gaussian errors.

### Introduction



#### These assumptions have taken us a long way!



Hurricane Sandy (2012)



So, why bother trying to solve this problem?

- Model resolution: Gaussian assumption may be more severe at smaller scales.
- Adoption of ensemble data assimilation: useful information is ignored when ensembles are assumed to follow a Guassian error distribution.

Will our current data assimilation methods still be the best option as model resolution and ensemble size increase?



Let  $\mathbf{x}_t$  be a vector of state variables, and  $\mathbf{y}_t$  be a vector of observations that are valid at time t.

 $\mathbf{x}_t$  and  $\mathbf{y}_t$  are given by the (possibly) nonlinear equations:

$$egin{array}{rcl} \mathbf{x}_t &=& M(\mathbf{x}_{t-1}) + \eta_t, \ \mathbf{y}_t &=& H(\mathbf{x}_t) + \epsilon_t, \end{array}$$

where  $\eta_t$  and  $\epsilon_t$  are stochastic terms that represent model and observation errors.

Let  $y_t$  denote all observations up to time t.



The probability of  $\mathbf{x}_t$ , given all information up to this time is given by Bayes' theorem:

$$p(\mathbf{x}_t|y_t) = rac{p(\mathbf{y}_t|\mathbf{x}_t)p(\mathbf{x}_t|y_{t-1})}{\int p(\mathbf{y}_t|\mathbf{x}_t)p(\mathbf{x}_t|y_{t-1})d\mathbf{x}_t}.$$

**Posterior:**  $p(\mathbf{x}_t|y_t)$ 

**Likelihood:**  $p(\mathbf{y}_t | \mathbf{x}_t)$ 

**Prior:**  $p(\mathbf{x}_t|y_{t-1})$ 

**Evidence/support:**  $\int p(\mathbf{y}_t | \mathbf{x}_t) p(\mathbf{x}_t | y_{t-1})$ 



Use an ensemble to construct an emperical estimate of the posterior probability distribution:

$$p(\mathbf{x}_t|y_t) \approx \sum_{n=1}^N w_t^n \delta(\mathbf{x}_t - \mathbf{x}_t^n).$$

Use the particle representation of  $p(\mathbf{x}_t|y_t)$  to approximate expectations of functions f of the model state:

$$\overline{f(\mathbf{x}_t)} = \int f(\mathbf{x}_t) p(\mathbf{x}_t | y_t) d\mathbf{x}_t,$$
$$\approx \sum_{n=1}^N w_t^n f(\mathbf{x}_t^n).$$

Examples of  $\overline{f(\mathbf{x}_t)}$  are the mean and covariance of  $\mathbf{x}_t$ .



#### For the simplest particle filter, the weights are given by

$$w_t^n \propto w_{t-1}^n p(\mathbf{y}_t | \mathbf{x}_t^n)$$



When the particles begin to move away from observations, a resampling step is needed to produce a more informative ensemble.

The simplest example is the **bootstrap filter**: sample *N* new particles from the posterior error distribution with replacement, and assign each new particle a weight of  $\frac{1}{N}$ . The expression for  $w_t^n$  at the next cycle simplifies further:

$$w_t^n = \frac{p(\mathbf{y}_t|\mathbf{x}_t^n)}{\sum_{n=1}^N (\mathbf{y}_t|\mathbf{x}_t^n)}$$



For Gaussian observation errors, the  $n^{th}$  weight is given by

$$w_t^n = A \frac{\exp\left\{-\frac{1}{2}[\mathbf{y}_t - H(\mathbf{x}_t^n)]^T \mathbf{R}^{-1}[\mathbf{y}_t - H(\mathbf{x}_t^n)]\right\}}{\sum_{k=1}^N \exp\left\{-\frac{1}{2}[\mathbf{y}_t - H(\mathbf{x}_t^k)]^T \mathbf{R}^{-1}[\mathbf{y}_t - H(\mathbf{x}_t^k)]\right\}},$$

and the posterior mean can be approximated with

$$\overline{\mathbf{x}_t} \approx \sum_{n=1}^N w_t^n \mathbf{x}_t^n$$



- Filter degeneracy: a finite ensemble will eventually lose track of the signal, in which case, the weights become concentrated on a small number of particles. This problem occurs faster when the dimensions of x and y are large (Snyder et al. 2008, MWR).
- Resampling: the process of resampling from the posterior can be problematic, since duplicate particles will be produced. This is a larger problem for deterministic models; i.e., no stochastic terms.



# For high-dimensional systems, these weights are determined by applying a method called **sequential importance sampling**.

Particles are sampled from a subspace of the model space using a **proposal distribution** that is typically conditioned on observations between cycles (e.g., Leeuwen 2010, QJRMS).



Expand  $w_n^t$  from a scalar to a vector with the same dimension as **x**, so that

$$\overline{\mathbf{x}_t} \approx \sum_{n=1}^N w_t^n \circ \mathbf{x}_t^n.$$

This is the same as estimating a weight for every state variable in each single particle.



One way of obtaining this result is to evaluate the likelihood locally for each variable, so that observations at large distances from a grid point will not effect the local weight.

This requires a likelihood function that depends on the physical distance between the model grid point and observation locations.

A simple example is to use a function that decays exponentially away from the location of an observation, so that the likelihood of the the  $i^{th}$  variable given the  $j^{th}$ observation is written

$$p(y_{t,j}|x_{t,i}^n) = [p(y_{t,j}|\mathbf{x}_t^n) - \frac{1}{N_eN_y}]exp(-\frac{d_{i,j}}{R}) + \frac{1}{N_eN_y},$$

where  $d_{i,j}$  is the physical distance between the observation and model grid point, and R is a tunable localization radius.



## Local likelihood particle filter (LLPF)

The Lorenze-96 model is used to test this idea and compare with  $\mathsf{EnKF}$ :

$$\frac{dx_{t+1,i}}{dt} = (x_{t,i+1} - x_{t,i-1})x_{t,i-1} - x_{t,i} + 10.$$

Experiment details:

- 100 variables
- Observations are taken from a "truth" run at every other grid point, with added Gaussian noise ( $\sigma = 1$ ).
- These observations are assimilated every 12 h (dt = 0.1) for 1000 cycles.
- Relaxation coefficient and localization radius for EnKF are tuned for each ensemble size.

## Local likelihood particle filter (LLPF)

$$N_e = 50$$

PENN<u>State</u>



RMSEs are averaged every 30 cycles.



 $N_e = 100$ 

PENNSTATE



RMSEs are averaged every 30 cycles.



$$N_{e} = 200$$

PENN<u>State</u>



RMSEs are averaged every 30 cycles.



- A new approach to particle filtering, called LLPF, has been introduced for systems that contain a large spatial dimension.
- This method makes use of a distance-dependent likelihood function to prevent filter degeneracy.
- The LLPF is shown to be stable without the use of an optimal proposal density.
- Much more work needs to be done; e.g., resampling, inflation, better choices for likelihood function, etc.